Currents in complex polymers: an example of superstatistics for short time series

G. Cigdem Yalcin\textsuperscript{1*} and Christian Beck\textsuperscript{2}

\textsuperscript{1}Department of Physics, Istanbul University, 34134, Vezneciler, Istanbul, Turkey
\textsuperscript{2}School of Mathematical Sciences, Queen Mary, University of London, London E1 4NS, UK

We apply superstatistical techniques to an experimental time series of measured transient currents through a thin Aluminium-PMMA-Aluminium film. We show that in good approximation the current can be approximated by local Gaussian processes with fluctuating variance. The marginal density exhibits 'fat tails' and is well modelled by a superstatistical model. Our techniques can be generally be applied to other short time series as well.

I. INTRODUCTION

Many time series generated by complex systems in nature are superstatistical, i.e. they consist of a superposition of several dynamics on well-separated time scales. Often, there is locally a simple dynamics (for example, a Gaussian process) and the parameters of that simple process fluctuate on a much larger scale. Such varying parameters describe a changing environment of the local system under consideration. Often the relevant measured time series consists locally of a Gaussian process, with the variance of those Gaussians evolving on a longer time scale. In nonequilibrium statistical mechanics, the technique of superstatistics was introduced in \cite{1} and has since then provided a powerful tool to describe a large variety of complex systems for which there is change of environmental conditions \cite{2–9}. A superstatistical complex system is mathematically described as a multi-scale superposition of two (or several) statistics, one corresponding to local equilibrium statistical mechanics (on a mesoscopic level modeled by a linear Langevin equation leading to locally Gaussian behavior) and the other one corresponding to a slowly varying parameter $a$ of the system. Essential for this approach is the fact that there is sufficient time scale separation, i.e. the local relaxation time of the system must be much shorter than the typical time scale on which the parameter $a$ changes. There is interesting mathematics associated with superstatistical complex systems. For example, in a recent paper \cite{10} Hanel, Thurner and Gell-Mann showed that the superstatistical distribution function $f(a)$ cannot change in an arbitrary way under macroscopic state changes of the system, and that the underlying transformation group is the Euclidean group in 1 dimension.

In most applications in nonequilibrium statistical mechanics the slowly varying control parameter $a$, underlying the superstatistical dynamics, is the local inverse temperature $\beta$ of the system, i.e. $a = \beta$. However, in general the control parameter $a$ can also have a different meaning. For example in mathematical finance, $a$ is a volatility parameter describing changing market behavior and turbulences in the market. For a time series, $a$ is usually a local variance parameter that can be directly extracted from suitably chosen time slices. There are numerous interesting applications of the superstatistics concept to real-world problems, for example to train delay statistics\cite{11}, hydrodynamic turbulence\cite{12} and cancer survival statistics\cite{13}. For further applications, see \cite{14–19}.

In this paper, we apply for the first time superstatistical techniques to a complex systems of solid state physics. We are interested in the statistical properties of currents flowing through thin films when a voltage is applied. Our measured system is a thin aluminium-polymethylmethacrylate-aluminium film (Al-PMMA-Al). A detailed experimental investigation of this system was presented in \cite{20, 21}. A particular feature of this system is that when a voltage is applied, the corresponding measured currents exhibit transient (nonstationary) behavior with strong fluctuations. This can be formally associated with transient chaotic behavior, and in the above papers formally a small positive Lyapunov exponent was extracted from the data. More recently, Yalcin et al \cite{22, 23} made a statistical analysis of the fluctuating current data using $q$-statistics, inspired by the fact that $q$-statistics \cite{24} is often used for weakly chaotic systems which have approximately zero Lyapunov exponent \cite{25, 27}.

In this paper we show that the measured current data can be well modelled by a superstatistical dynamics. However, in contrast to previous superstatistical time series analysis techniques described in \cite{9}, one has a typical problem for these kinds of data: The time series is not long enough to apply the standard superstatistical techniques developed in \cite{9}. Still, in this paper we show what one can typically do for a short time series. Our method described in the following is quite simple and general and can be applied to other small data sets in a similar way.

* corresponding author, phone +90 212 455 57 00 ext 15489, Fax +90 212 455 5855, gcyalcin@istanbul.edu.tr
II. ANALYZING THE EXPERIMENTAL TIME SERIES

Figs. 1-3 show typical time series of increments $I(t) = i(t+1) - i(t)$ of the measured current $i(t)$ through the Al-PMMA-Al film at three different temperatures. The average value has been subtracted, and all data are rescaled by dividing through the standard deviation $\sigma$ determined from the entire time series. We define $u(t) = (I(t) - \langle I(t) \rangle) / \sigma$.

One clearly observes regions of strong activity interwoven with periods of much calmer behavior, similar as for turbulent flows of share price evolution data. To analyse these data, we divide the time series into a couple of windows with qualitatively similar behavior. These windows are shown as vertical lines in Fig.1-3.

The choice of the windows is rather arbitrary, they should be large enough to contain enough statistics, but still be significantly shorter than the entire time series.

For each local window, we look at a histogram of the fluctuating currents. As shown in Fig. 4-7, locally the behavior
FIG. 3: The current magnitude differences $I(t) = i(t+1) - i(t)$ for the transient current through thin Al-PMMA-Al film at $T = 313K$.

FIG. 4: Gaussian fit to the observed local probability density for data points 1-300 of the measured time series of current differences $I(t)$ at $T = 303K$.

is well approximated by a Gaussian distribution

$$p(u|\beta) = \sqrt{\frac{\beta}{2\pi}} e^{\frac{-1}{2} \beta u^2}$$

However, the local variance parameter $\beta$ varies strongly from window to window and is thus itself a random variable, as expected for superstatistical systems.

Tab. 1 shows the values of extracted variance parameters $\beta_i$ for the various windows and for different temperatures $T$ where the experiment was performed.

We can only extract a few values $\beta_i$ from our short time series. The general idea of superstatistics is that the marginal distribution $p(u)$ is given as a superposition

$$p(u) = \int f(\beta)p(u|\beta)d\beta$$
for a suitable $f(\beta)$ describing the distribution of $\beta$-values. For a short time series, as in our case, one can only make guesses of the relevant $f(\beta)$. For this it is useful to look at the histogram of the entire time series, sampling up the behavior in all windows. This is shown in Fig. 8.

Apparently, the histogram data of the entire time series are well-fitted by $q$-Gaussians of the form

$$p(u) \sim (1 + \frac{1}{2} \tilde{\beta}(q - 1)u^2)^{\frac{-1}{q-1}}.$$  

The corresponding parameters $\tilde{\beta}$ and $q$ are shown in Tab.1.

We cannot, as for longer time series [9], extract the relevant distribution $f(\beta)$ from a histogram of $\beta_i$-values, since there are only 4 data points. But the fitting result [9] can be used to theoretically predict $f(\beta)$. In our case $f(\beta)$ is in good approximation a $\chi^2$ (or $\Gamma$)-distribution, meaning that if we were able to look at much longer time series, then the $\beta_i$ would be distributed according to

$$f(\beta) = \frac{1}{\Gamma(n/2)} \left\{ \frac{n}{2\beta_0} \right\}^{n/2} \beta^{n/2-1} e^{-\beta n/2} \left\{ -\frac{n\beta}{2\beta_0} \right\}$$  

FIG. 5: Same as Fig. 4 for data points 301-600 of the measured time series

FIG. 6: Same as Fig. 4 for data points 601-800 of the measured time series
This follows from doing the integral (2) using the distribution (4), which directly leads to the observed fitting result (3). The relation between the parameters \( n \) and \( q \) is

\[
q = 1 + \frac{2}{n + 1}. \tag{5}
\]

<table>
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<th>( T )</th>
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<th>( \beta_2 )</th>
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<td>22</td>
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</table>

TABLE I: The values of extracted variance parameters \( \beta_1, \beta_2, \beta_3, \beta_4 \), and \( \beta_0 \) for the various slices, for the measured current through thin Al-PMMA-Al film at temperature \( T = 295, 303 \) and \( 313K \). Also shown \( q \) and \( \tilde{\beta} \).
Moreover, the parameter $\tilde{\beta}$ is related to the average value of $\beta$ by 

$$\beta_0 := \int f(\beta)\beta d\beta = (3 - q) / 2$$

The predicted superstatistical distributions $f(\beta)$ are plotted in Fig. 9 for the three different temperatures. We can also calculate the average value of $\beta_0$ directly from the time series, namely from the values $\beta_i$ fitted in the 4 windows, i.e. $\beta_0 = 1/4(\beta_1 + \beta_2 + \beta_3 + \beta_4)$. The result is shown in Tab. 1 as well. This sum has huge error terms, because only 4 windows enter, and the partitioning into windows is rather arbitrary. Nevertheless, the obtained average values are consistent with the theoretically predicted result (6) within the statistical error bounds.

To summarize, for very long time series, one has theoretical methods to decide whether a given time series is superstatistical or not [9]. For short time series, as in our case, these methods are not applicable. Still one can do the method proposed in the current paper quite generally. First one looks whether a partitioning into a few windows is possible such that locally Gaussian behavior is observed. Then one determines the local variance parameters $\beta_i$. From a histogram of the entire time series one obtains a guess what the relevant distribution $f(\beta)$ could be—in our case a $\chi^2$-distribution but in general these can be other distributions such as e.g. the lognormal or inverse-$\chi^2$ distribution [2]. After that guess, one can check consistency of derived parameter relations within the statistical error bounds, in our case given by relation (6) but in general given by other relations, depending on $f(\beta)$.

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