

Adaptive Dose-Finding for Time-to-Event Outcomes with Adaptive Choice of Patient Number Based on Response Rate

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Outline



- 2 Accelerated Failure Time (AFT) Model
- Optimal Design (OD) Problem
- 4 Adaptive Designs (AD)





Outline



2 Accelerated Failure Time (AFT) Model

③ Optimal Design (OD) Problem

4 Adaptive Designs (AD)





Objectives

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- O To determine the requisite sample size based on some predefined stopping criterion.
- O To determine doses needed for dose-response estimation and allocation proportions at these doses by adapting to available data.



Main Results

• An adaptive design which may improve efficiency of a dose-finding clinical trial with TTE outcomes has been obtained.



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- A stopping criterion which provides an adaptive choice of a sample size has been proposed:
 - the required sample size is smaller when more events are observed in the trial; the sample size increases when the amount of censored data increases.



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2 Accelerated Failure Time (AFT) Model

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5 Simulations



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$$\log T = \beta_0 + \beta_1 x + \beta_2 x^2 + b\varepsilon,$$



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$$\log T = \beta_0 + \beta_1 x + \beta_2 x^2 + b\varepsilon,$$

where

- x corresponds to a dose (treatment arm),
- shape parameter: $\lambda = \exp(\beta_0 + \beta_1 x + \beta_2 x^2)$,
- scale parameter: $p = b^{-1}$,
- and $\varepsilon \sim f_{\varepsilon}(w) = exp(w exp(w)) extreme value distribution.$



Dose-response relationship: $E(T|x) = \exp(\beta_0 + \beta_1 x + \beta_2 x^2)\Gamma(1+b)$.



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Censoring





Likelihood and Fisher Information

• For a sample of N patients and vector of parameters $\boldsymbol{\theta} = (\boldsymbol{\beta}^{\mathrm{T}}, b)^{\mathrm{T}}$ one can calculate log-*likelihood function* log $\mathcal{L}(\boldsymbol{\theta})$.



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- Then, MLEs of unknown model parameters $(\hat{\theta}_{MLE})$ are the solutions of score equations

$$\frac{\partial \mathrm{log}\,\mathcal{L}(\boldsymbol{\theta})}{\partial \boldsymbol{\theta}} = \left(\begin{array}{c} \frac{\partial \mathrm{log}\,\mathcal{L}(\boldsymbol{\theta})}{\partial \boldsymbol{\beta}}\\ \frac{\partial \mathrm{log}\,\mathcal{L}(\boldsymbol{\theta})}{\partial \boldsymbol{b}} \end{array}\right) = \mathbf{0}$$



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ight) = oldsymbol{0}$$

• The corresponding Fisher Information Matrix is

$$I(\boldsymbol{\theta}) = -\mathbf{E}\left(\frac{\partial^2 \log \mathcal{L}(\boldsymbol{\theta})}{\partial \boldsymbol{\theta} \partial \boldsymbol{\theta}^{\mathrm{T}}}\right)$$



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$$\xi = \left(\begin{array}{cccc} x_1 & x_2 & \dots & x_K \\ w_1 & w_2 & \dots & w_K \end{array}\right),$$



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- K is a number of doses (treatment arms).
- $x_k, k = 1, \ldots K$ a set of selected doses.
- $w_k, k = 1, ..., K$ are the proportions of patients assigned to corresponding doses.



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$$x_k \in \mathcal{X} = [-1; 1], \qquad \sum_{k=1}^K w_k = 1.$$



• For a given design ξ the full *Fisher Information Matrix* is

$$FIM(\xi, \boldsymbol{\theta}) = N \sum_{k=1}^{K} w_k I(\boldsymbol{\theta}|x_k).$$



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$$FIM(\xi, \boldsymbol{\theta}) = N \sum_{k=1}^{K} w_k I(\boldsymbol{\theta} | x_k).$$

• Then, a *D-optimal design* is determined as a solution of the following optimization problem

$$\xi_D^* = \arg\max_{\xi} |FIM(\xi, \theta)|.$$



• Without censoring, *D*-optimal design is a 3-points balanced (uniform) design^a

$$\left(\begin{array}{ccc} -1 & 0 & 1 \\ 1/3 & 1/3 & 1/3 \end{array}
ight),$$

where
$$\begin{bmatrix} -1 & - & \text{minimum dose} \\ 0 & - & \text{average dose} \\ 1 & - & \text{maximum dose} \end{bmatrix}$$

 $[^]a\mathrm{Ryeznik}$ Y, Hooker AC, Sverdlov O Adaptive designs for dose finding clinical trials with time-to-event outcomes. PAGE 24 (2015) Abstr 3608 [www.page-meeting.org/?abstract=3608]



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• In the presence of censoring *D*-optimal design still has 3 points but it is shifted from the uniform design^a.

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D-efficiency: $D_{eff} = \left(\frac{|FIM(\xi_D^*, \theta)|}{|FIM(\xi_U^*, \theta)|}\right)^{\frac{1}{4}}$



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- D-optimal design ξ depends on a model parameters θ which are *unknown* at the beginning of a study.
- In order to address this issue *adaptive design* is proosed.

















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Updated Design

Two types of design updating are possible:

either
$$\tilde{\xi}_D = \arg \max_{\xi} \left\{ \log |FIM(\xi, \hat{\theta}_{MLE})| \right\}$$

or
$$\widetilde{\xi}_{ED} = \arg \max_{\xi} \int_{\Theta} \log |FIM(\xi, \theta)| \widehat{\pi}(\theta) d\theta$$
,

where

 $\widehat{\boldsymbol{\theta}}_{MLE}$ are maximum likelihood estimators of $\boldsymbol{\theta}$, $\widehat{\pi}(\boldsymbol{\theta})$ is a posterior distribution of $\boldsymbol{\theta}$

given data.



Stopping Criterion

The following rule is proposed as a stopping criterion:

$$\max\left\{\frac{SD_{\widehat{\beta}_0}}{|\widehat{\beta}_0|}, \quad \frac{SD_{\widehat{\beta}_1}}{|\widehat{\beta}_1|}, \quad \frac{SD_{\widehat{\beta}_2}}{|\widehat{\beta}_2|}, \quad \frac{SD_{\widehat{b}}}{|\widehat{b}|}\right\} \leq \alpha, \qquad \alpha > 0,$$

where

$$\begin{split} & \left(\widehat{\beta}_{0}, \widehat{\beta}_{1}, \widehat{\beta}_{2}, \widehat{b}\right) = \widehat{\boldsymbol{\theta}}_{MLE} \\ & \left(SD_{\widehat{\beta}_{0}}, SD_{\widehat{\beta}_{1}}, SD_{\widehat{\beta}_{2}}, SD_{\widehat{b}}\right) = \operatorname{diag}\left\{FIM_{obs}^{-1}(\widehat{\boldsymbol{\theta}}_{MLE})\right\} \end{split}$$



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Simulations Setup

- For 3 different scales and 4 different shapes we simulate 12 models.
- We consider 3 possible values of a response rate¹
- For each of 12 scenarios number of simulation $n_{sim} = 1000$.
- The number of subject in cohort is 30.
- Parameter α for the stoping criterion is 0.25.
- We stop simulations if number of randomized patients achives 2100 but stopping criterion is not satisfied.

¹By response rate we assume proportion of uncensored observations.





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Dose-Response Relationship (Response Rate ~ 20%)







Boxplots of Number of Patients (Response Rate ~ 20%)

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UNIVERSITET 2-Stage AD (Bayesian update vs. MLEs)

When first cohort is randomized we do just one design update:

either
$$\tilde{\xi}_D^{(2)} = \arg \max_{\xi} \left\{ \log |FIM(\xi, \hat{\theta}_{MLE})| \right\}$$

or $\tilde{\xi}_{ED}^{(2)} = \arg \max_{\xi} \int_{\Theta} \log |FIM(\xi, \theta)| \hat{\pi}(\theta) d\theta,$

where

 $\widehat{\boldsymbol{\theta}}_{MLE}$ are maximum likelihood estimators of $\boldsymbol{\theta},$

$$\widehat{\pi}(\boldsymbol{\theta})$$
 is a posterior distribution of $\boldsymbol{\theta}$

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²-Stage AD (Bayesian update vs. MLEs)

$$RelD_{eff} = \left(|FIM(\widetilde{\xi}, \boldsymbol{\theta})| / |FIM(\xi_D^*, \boldsymbol{\theta})| \right)^{\frac{1}{4}}, \quad \boldsymbol{\theta} \text{ is TRUE}$$



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2-Stage AD (Bayesian update vs. MLEs)

Boxplots of Number of Patients (Response Rate ~ 50%)





2-Stage AD (Bayesian update vs. MLEs)



UNVERSITET 2-Stage AD (Bayesian update vs. MLEs)



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Summary

• The proposed adaptive design may improve efficiency of a dose-finding clinical trial with censored TTE outcomes. It allows amendation the dose levels and allocation proportions at these doses for the next cohorts of patients after interim analysis based on available data.



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- The proposed adaptive design may improve efficiency of a dose-finding clinical trial with censored TTE outcomes. It allows amendation the dose levels and allocation proportions at these doses for the next cohorts of patients after interim analysis based on available data.
- On the proposed stoping criterion allows adaptive choice of a requisite sample size. For high response rate we need fewer patients, while the number increases for a low response rate.



Summary

- The proposed adaptive design may improve efficiency of a dose-finding clinical trial with censored TTE outcomes. It allows amendation the dose levels and allocation proportions at these doses for the next cohorts of patients after interim analysis based on available data.
- On the proposed stoping criterion allows adaptive choice of a requisite sample size. For high response rate we need fewer patients, while the number increases for a low response rate.
- It seems that adaptive designs with Bayesian updating outperform adaptive designs based on MLE updating.



Thank You!

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