

Experiment Design Based on Bayes Risk and Weighted Bayes Risk with Application to Pharmacokinetic Systems

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Introduction

- **Multiple Model Optimal (MMOpt) Design**
 - Captures essential elements of Bayesian Experiment Design without the excessive computation
 - Minimizes a recent theoretical overbound on the Bayes Risk (Blackmore et. al. 2008 [4])
 - Intended for multiple model (MM) parametrizations which form the basis of the USC BestDose software (corresponds to the support points in a nonparametric population model)
 - Has several advantages relative to D-optimality and other criteria based on the asymptotic Fisher Information matrix for nonlinear problems
- **Contribution of present paper, since the last PODE, is to generalize MMOpt by introducing a weighting into the Bayes Risk Cost**
 - New result shows that simple analytical overbound of [4] is preserved in the weighted case
 - Weights allow MMOpt experiment design to address many problems of practical interest (AUC estimation, what best future dose to give, etc.)
- **Numerical examples demonstrate MMOpt on several relevant PK problems**

- **Dynamic Model and Measurements**

$\dot{x}(t) = f(x(t), d(t), \theta)$ State x , Input d , Parameter $\theta \in R^p$

$\eta_k = h(x(t_k), \theta)$, System output at time t_k

$y_k = \eta_k + \sigma_k n_k$, Noisy measurement at time t_k

$n_k \sim N(0, 1)$, Gaussian measurement noise

$\xi = \{t_1, \dots, t_n\}$, Experiment design (optimal sampling)

- **D-Optimal Design**

$$\max_{\xi} |M|$$

where the Fisher Information Matrix M is given by,

$$M(\theta, \xi) = \sum_{k=1}^n \frac{1}{\sigma_k^2} \left[\frac{\partial \eta_k}{\partial \theta} \frac{\partial \eta_k}{\partial \theta^T} \right] \Bigg|_{\theta=\bar{\theta}}$$

- Herein, $M(\theta, \xi)$ is assumed to be a function of θ
(i.e., nonlinear problems)

D-Optimal Design for Nonlinear Problems

- **D-optimality** (traditional) maximizes the determinant of the Fisher Information Matrix (Fedorov 1972 [20], Silvey 1980 [19])
 - $\max |M|$, where M is Fisher Information Matrix, and $|(\cdot)|$ is determinant
 - Useful tool has become standard in design of clinical experiments
- For nonlinear problems, **MMOpt** offers several advantages relative to D-optimality and other criteria based on the asymptotic Fisher Information matrix
 - Avoids **circular reasoning** associated with having to know a patient's true parameters in order to design an experiment
 - Avoids using an **asymptotic information measure** when placing only a small number of samples
- **To robustify D-optimal design**, an expectation is taken with respect to certain functions of prior information giving rise to ED, EID, and ELD (or API) optimal designs
 - Chaloner [13], Pronzato [14][15], Tod [16], D'Argenio [17]

Definition of ED, EID, API

- Robust D-Optimal Designs

$$\text{ED: } \arg \max_{\xi} E_{\theta} (|M|)$$

$$\text{EID: } \arg \min_{\xi} E_{\theta} \left(\frac{1}{|M|} \right)$$

$$\text{API: } \arg \max_{\xi} E_{\theta} (\log |M|)$$

where,

$\theta \in R^p$ - Parameter Vector

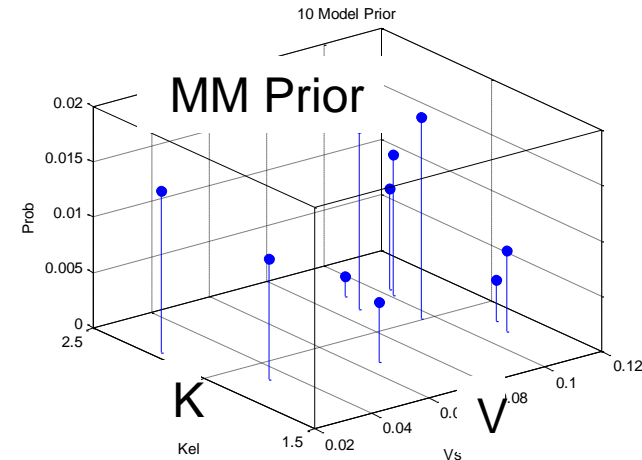
$\eta = \{t_1, \dots, t_n\}$ - Experiment design

M - Fisher Information Matrix

- All above design metrics require Fisher Matrix M to be nonsingular, and hence **require at least p samples to be taken**, where $p = \#$ parameters

Multiple Model Optimal Design

- USC *BestDose* software [3] is based on a multiple model (MM) parametrization of the Bayesian prior (i.e., discrete support points in the nonparametric population model)
 - Nonparametric Maximum Likelihood (NPML) estimation of a population model has the form of a MM prior (Mallet [5], Lindsay [6]).
 - Software for population NPML modeling is available, e.g., NPEM (Schumitzky [7][11]), NPAG (Leary [8], Baek [9]), USC*PACK (Jelliffe [10], and **Pmetrics** in Bestdose [3].
- Experiment design for MM (i.e., discrete) models is a subject found in classification theory
 - How do we sample the patient to find out which support point he best corresponds to?
 - Classifying patients is fundamentally different from trying to estimate patient's parameters
- Treating MM experiment design in the context of classification theory leads to the mathematical problem of minimizing Bayes risk (Duda et. al. [21])



Multiple Model Optimal Design (MMOpt)

- **Bayes Rule**

$$p(H_i|y, u) = \frac{p(y|H_i, u)p(H_i)}{p(y|u)}, \quad i = 1, \dots, m$$

H_i - Hypothesis that model i corresponds to true subject

u - Experiment design variable (to be optimized over)

- **Design Rule for MM Classifier**

If $p(H_j|y, u) = \max_i \{p(H_i|y, u)\}$, then

1. H_j is classified as TRUE
(i.e., j 'th model is classified as true subject)
2. H_i for $i \neq j$ is classified as FALSE

- **Design Regions**

MM classifier breaks y into m regions R_i , $i = 1, \dots, m$
such that H_j is classified as TRUE when $y \in R_j$.

Multiple Model Optimal Design (Cont'd)

- **Bayes Risk (i.e., Probability of MM Classifier Being Wrong)**

$$P(\text{error}) = \sum_i^m \sum_{j \neq i}^m P(y \in R_j, H_i | u)$$

(Sum of probabilities over all possible ways of making a mistake)

- **Bayes Risk represents a cost function to be minimized**

- Consistent with a Bayesian experiment design philosophy

- **Result: (Blackmore et. al. 2008)**

The Bayes Risk is upper bounded as follows:

$$P(\text{error}) \leq \sum_i^m \sum_{j > i}^m P(H_i)^{\frac{1}{2}} P(H_j)^{\frac{1}{2}} e^{-k(i,j)} \quad (1)$$

where,

$$k(i,j) = \frac{1}{4}(\mu(j) - \mu(i))^T \left(\Sigma(i) + \Sigma(j) \right)^{-1} (\mu(j) - \mu(i)) + \frac{1}{2} \ln \frac{|\frac{1}{2}(\Sigma(i) + \Sigma(j))|}{\sqrt{|\Sigma(i)||\Sigma(j)|}}$$

- **MMOpt minimizes upper bound (1) on the probability that the true subject will be incorrectly classified**

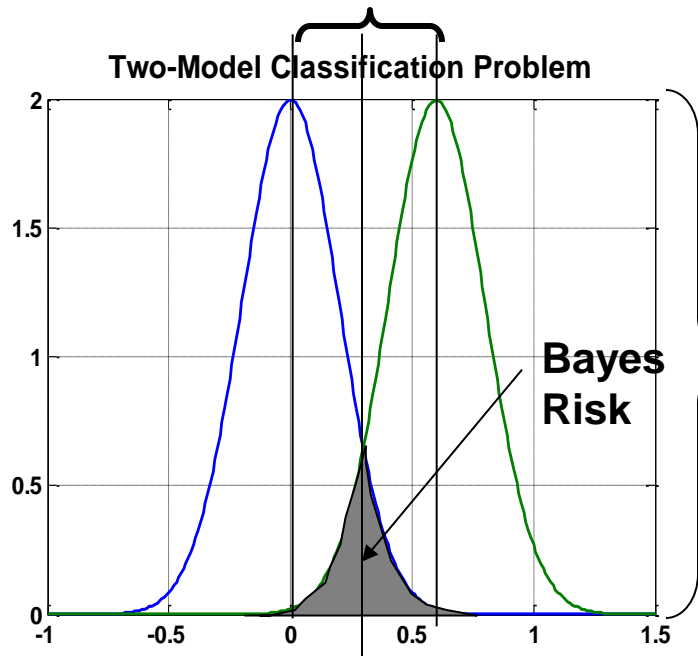
Model Response Separation $r(t)$

- Model Response Separation $r(t)$ is the separation between two model responses at a given time t

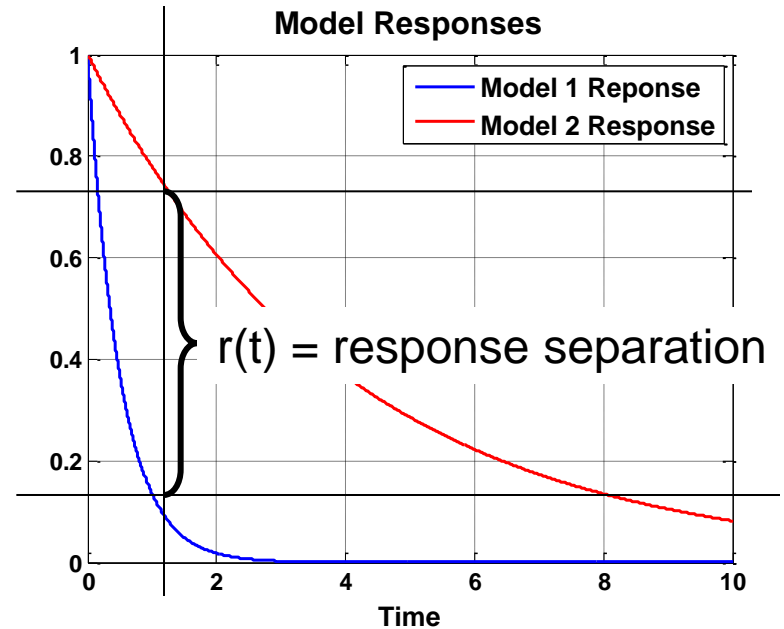
$$r(t) = |\eta(t, a_1) - \eta(t, a_2)|$$

- Defines natural statistic for discriminating between two models
- Bayes Risk is shown in gray area

$r(t)$ =response separation



Pull Gaussians apart to minimize gray area



- Bayes Risk (gray area) decreases as response separation $r(t)$ increases
- Models are best discriminated by sampling at a time t that maximizes $r(t)$

MMOpt Example: 4-Models (1/2)

- Two-Parameters a, b

$$y_i = \eta(t_i, a, b) + \sigma n_i$$

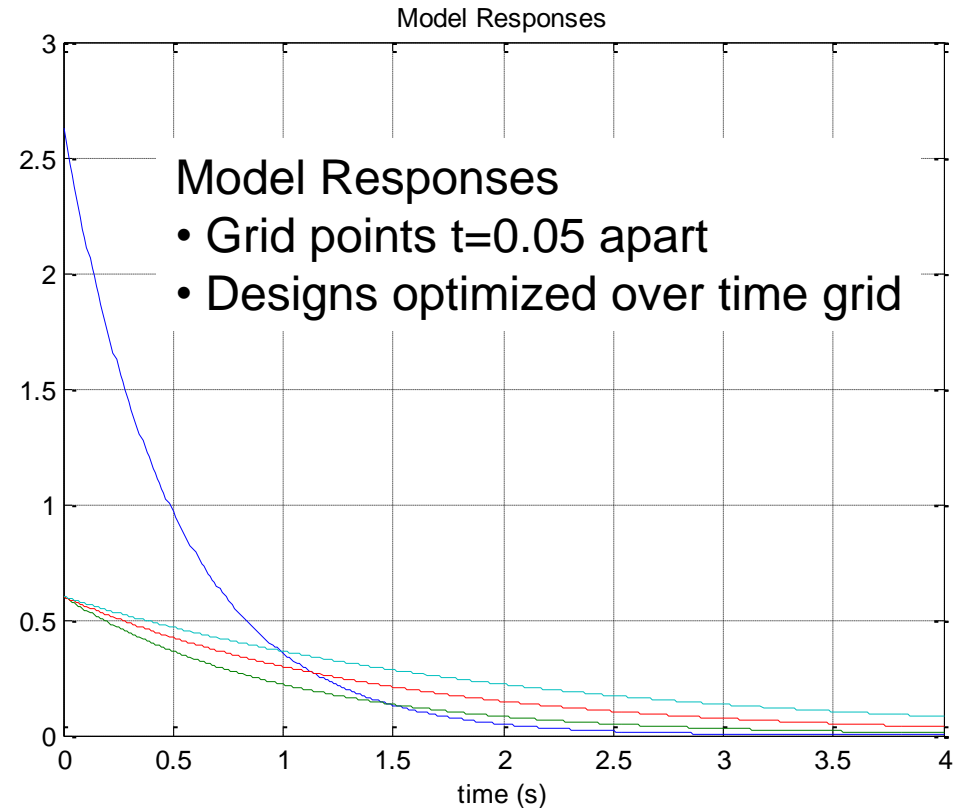
$$\eta(t, a, b) = be^{-at}$$

$$n_i \sim N(0, \sigma^2)$$

$$\sigma = 0.1$$

- Prior: $p_i = .25, \quad i = 1, \dots, 4$

Model Parameters		
#	a	b
1	2	2.625
2	1	0.6
3	0.7	0.6
4	0.5	0.6



MMOpt Example: 4-Models (2/2)

Design Metric	2-Sample Times		Bayes Risk	Bayes Risk 99%Conf *
MMOpt	0.45	1.4	0.32839	+/- 0.00070
ED	0	0.8	0.37028	+/- 0.00070
EID	0	1	0.36044	+/- 0.00072
API	0	0.95	0.36234	+/- 0.00072

Design Metric	3-Sample Times			Bayes Risk	Bayes Risk 99% Conf *
MMOpt	0.45	1.4	1.4	0.28065	+/- 0.00067
ED	0	0.7	0.9	0.32048	+/- 0.00067
EID	0	0	1	0.36034	+/- 0.00072
API	0	0.85	0.105	0.3099	+/- 0.00069

- MMOpt has smallest Bayes Risk of all designs studied

* evaluated based on Monte Carlo analysis using 1,400,000 runs per estimate

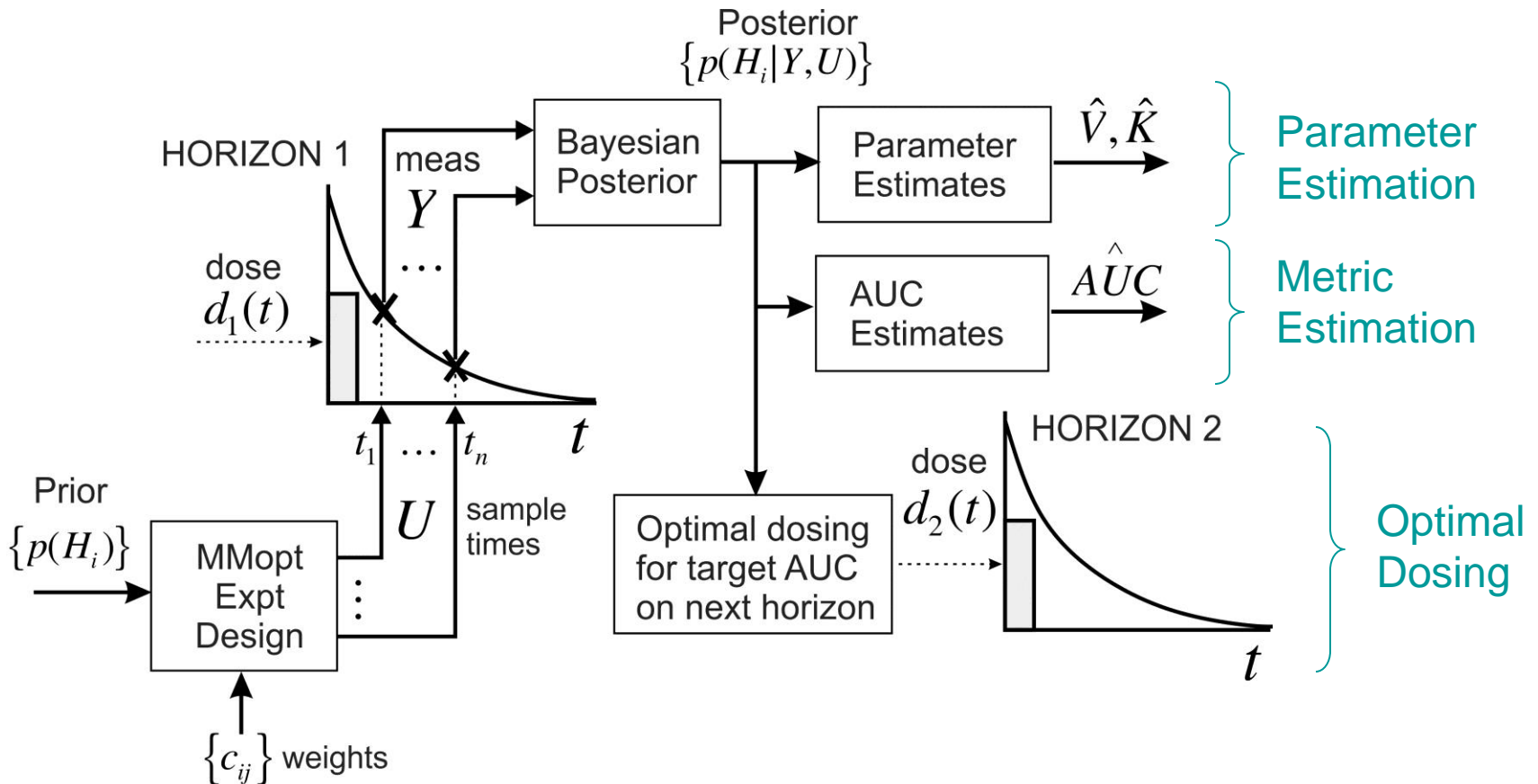
Comparison Table

	ED	EID	API	MMOpt
Invariant under regular <u>linear</u> reparametrization*	Yes	Yes	Yes	Yes
Invariant under regular <u>nonlinear</u> reparametrization*	No	No	Yes	Yes
Allows taking fewer than p samples, $p = \#$ of parameters	No	No	No	Yes
Can handle heterogeneous model structures	No	No	No	Yes
Gives known optimal solution to 2-model example*	No	No	No	Yes
Captures main elements of minimizing Bayes risk	No	No	No	Yes

* Proved in Bayard et. al., PODE 2013 [23]

Weighted MMOpt

- Introduce weights $\{c_{ij}\}$ to specify a cost for each type of classification error
- Assign c_{ij} as the cost of mistaking truth subject i for subject j ($j \neq i$)
- Choice of weights tailors experiment design to desired applications of interest



Weighted MMOpt

- **Weighted Bayes Risk (i.e., Expected Cost of MM Classifier Being Wrong)**

$$C(\text{error}) = \sum_i^m \sum_{j \neq i}^m c_{ij} P(y \in R_j, H_i | u)$$

(Sum of costs over all possible ways of making a mistake)

Here, c_{ij} is the cost of mistaking subject i for subject j ($j \neq i$)

- **Useful Result (new)**

The Weighted Bayes Risk is upper bounded as follows:

$$C(\text{error}) \leq \sum_i^m \sum_{j \neq i}^m \bar{c}_{ij} P(H_i)^{\frac{1}{2}} P(H_j)^{\frac{1}{2}} e^{-k(i,j)} \quad (2)$$

where,

$$k(i,j) = \frac{1}{4} (\mu(j) - \mu(i))^T \left(\Sigma(i) + \Sigma(j) \right)^{-1} (\mu(j) - \mu(i)) + \frac{1}{2} \ln \frac{|\frac{1}{2}(\Sigma(i) + \Sigma(j))|}{\sqrt{|\Sigma(i)||\Sigma(j)|}}$$

$$\bar{c}_{ij} = \max(c_{ij}, c_{ji})$$

- Result allows weighted bound-optimal designs to be systematically calculated as in the unweighted MMopt case

- **Weighted MMOpt minimizes upper bound (2) on the expected cost associated with the true subject being incorrectly classified**

Applications of MMOpt

- **Three Numerical Examples**
 - PK Estimation (unweighted MMOpt)
 - AUC Estimation (weighted MMOpt)
 - AUC Control (weighted MMOpt)
- **Results will be compared to ED optimal design EDopt**
 - Also compared to Bayes optimal design Bopt when computationally feasible to do so

PK Population Model with 10 Multiple Model Points - First-Order PK Model

- First-Order Model

$$\dot{x} = -Kx + d$$

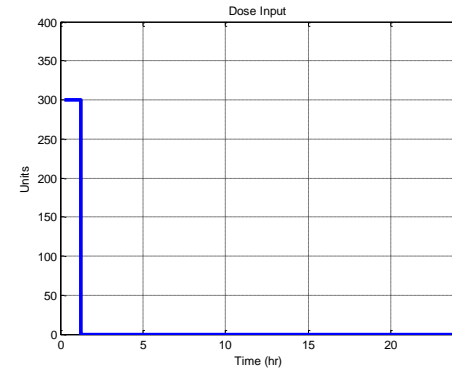
$$\eta_i = \frac{x(t_i)}{V}$$

$$y_i = \eta_i + \sigma_i n_i$$

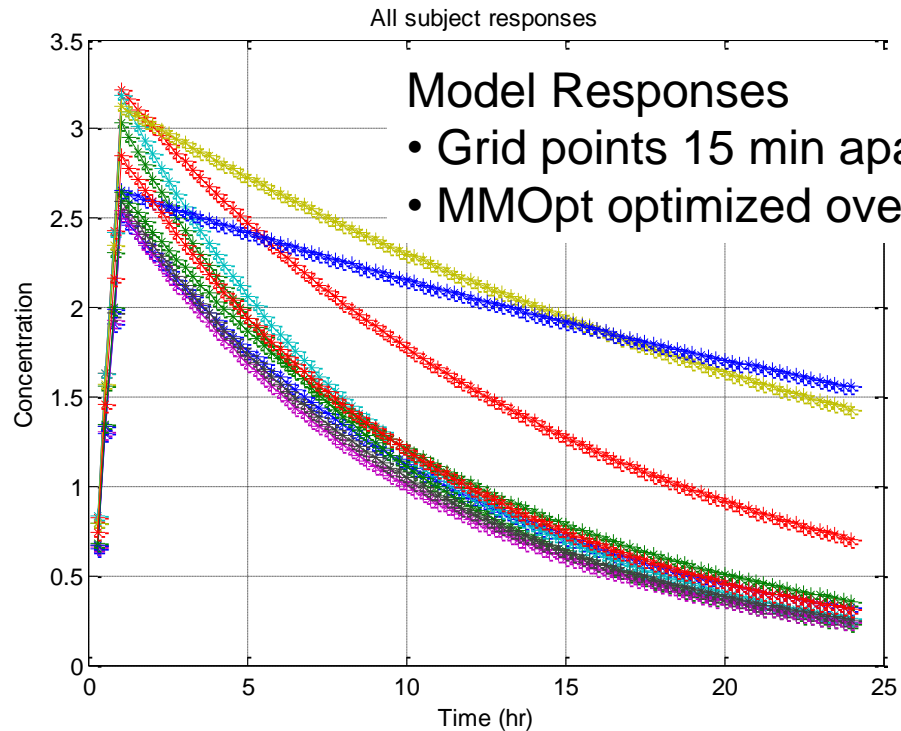
$$n_i \sim N(0, 1)$$

$$\sigma_i = 0.1$$

Dose input = 300 units for 1 hour, starting at time 0



Model Parameters		
#	K	V
1	0.090088	113.7451
2	0.111611	93.4326
3	0.066074	90.2832
4	0.108604	89.2334
5	0.103047	112.1093
6	0.033965	94.3847
7	0.100859	109.8633
8	0.023174	111.7920
9	0.087041	108.6670
10	0.095996	100.3418



Model Responses

- Grid points 15 min apart
- MMOpt optimized over time grid

Unweighted MMOpt for PK Estimation

- Summary of optimal 1,2 and 3 sample designs applied to PK Estimation

Design Metric	Samples (hr)			Bayes Risk (prob)	99% conf (prob)
	<i>1-Sample Design</i>				
B _{opt}	4.25			0.5474	±0.0015
MM _{opt}	4.25			0.5474	±0.0015
	<i>2-Sample Design</i>				
MM _{opt}	1	9.5		0.2947	±0.0014
ED _{opt}	1	24		0.3272	±0.0014
	<i>3-Sample Design</i>				
MM _{opt}	1	1	10.5	0.2325	±0.0013
ED _{opt}	1	1	24	0.2617	±0.0013

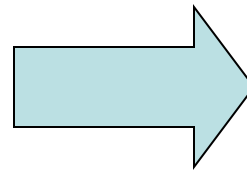
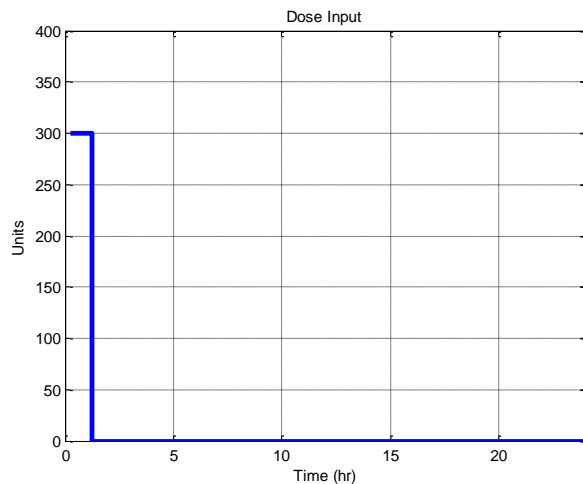
- 1 Sample Design: MM_{opt} performance equals Bayesian optimal design (both have Bayes Risk of 0.5474).**
- MM_{opt} performance improves on ED_{opt} design for 2 and 3 sample designs**
 - 2 Sample Design: Bayes Risk of 0.29 versus 0.33**
 - 3 Sample Design: Bayes Risk of 0.23 versus 0.26**
- All results are statistically significant to p<0.0001**

Weighted MMOpt for AUC Estimation

- **OBJECTIVE:** Design an experiment which is most informative about estimating patient's AUC
- In this case MMOpt weights are chosen as

$$c_{ij} = \left(\frac{D}{V_i K_i} - \frac{D}{V_j K_j} \right)^2$$

= Squared AUC error incurred if j 'th subject's AUC is used to estimate i 'th subject's AUC



#	AUC Responses
1	29.2767
2	28.7684
3	50.2902
4	30.9562
5	25.9683
6	93.5811
7	27.0741
8	115.8003
9	31.7176
10	31.1448
Mean	46.4578
STD	30.2314

Dose input $D = 300$ (300 units of drug infused over 1 hour)

AUC responses to dose D 18

Weighted MMOpt for AUC Estimation (Cont'd)

- Summary of optimal 1,2 and 3 sample designs applied to AUC estimation

Design Metric	Samples (hr)			RMS Error (AUC units)	99% conf (AUC units)
	<i>1-Sample Design</i>				
B _{opt} _C ₂	24			5.9059	±0.0270
MM _{opt} _C ₂	14			6.9789	±0.0265
MM _{opt}	4.25			21.6806	±0.0919
	<i>2-Sample Design</i>				
MM _{opt} _C ₂	1	13		1.8386	±0.0231
MM _{opt}	1	9.5		2.2346	±0.0483
ED _{opt}	1	24		2.2079	±0.0211
	<i>3-Sample Design</i>				
MM _{opt} _C ₂	1	10.25	10.25	1.4042	±0.0175
MM _{opt}	1	1	10.5	1.7025	±0.0382
ED _{opt}	1	1	24	1.8949	±0.0188

- 1 Sample Design: Weighted MMOpt performance approximates that of the Weighted Bayesian optimal design (RMS error of 6.98 versus 5.9 AUC units)**
- MMOpt performance improves on EDopt design**
 - 2 Sample Design: RMS error of 1.84 versus 2.21 (units of AUC)**
 - 3 Sample Design: RMS error of 1.40 versus 1.89 (units of AUC)**
- All results are statistically significant to p<0.0001**

Weighted MMOpt for AUC Control

- **OBJECTIVE:** Design an experiment most informative about next dose needed for patient to achieve a specified AUC of $\alpha_{des} = 40$
- In this case MMOpt weights are chosen as

$$c_{ij} = \left(\frac{D_j}{V_i K_i} - \alpha_{des} \right)^2$$

= Squared AUC error incurred if j 'th subject's ideal dose D_j is given to i 'th subject

#	Ideal Dose
1	409.8827
2	417.1242
3	238.6149
4	387.6442
5	462.1011
6	128.2311
7	443.2281
8	103.6267
9	378.3394
10	385.2965
Mean	335.4089
STD	35.8470

	$j = 1$	$j = 2$	$j = 3$	$j = 4$	$j = 5$	$j = 6$	$j = 7$	$j = 8$	$j = 9$	$j = 10$
$i = 1$	0	0.499	279	4.70	25.9	755	10.5	893	9.47	5.75
$i = 2$	0.482	0	293	7.99	18.6	767	6.26	903	13.8	9.31
$i = 3$	824	895	0	624	1403	342	1176	512	548	604
$i = 4$	5.26	9.25	236	0	59.0	716	32.8	858	0.921	0.0586
$i = 5$	20.4	15.1	374	41.5	0	835	2.66	962	52.5	44.2
$i = 6$	771	8121	1185	6548	10846	0	9654	58.9	6086	6430
$i = 7$	9.05	5.54	340	25.1	2.90	808	0	939	34.2	27.3
$i = 8$	13975	14643	2715	12019	19147	90.1	17184	0	11244	11821
$i = 9$	11.1	16.8	218	0.967	78.4	699	47.0	843	0	0.541
$i = 10$	6.51	10.9	231	0.0594	63.5	712	36.1	855	0.521	0

Ideal Doses $\{D_j\}$ to achieve desired AUC of $\alpha_{des} = 40$

Matrix of Weights $\{c_{ij}\}$

Weighted MMOpt for AUC Control (Cont'd)

- Summary of optimal 1,2 and 3 sample designs applied to AUC control

Design Metric	Samples (hr)			RMS Error (AUC units)	99% conf (AUC units)
	<i>1-Sample Design</i>				
Bopt_C ₁	12.5			3.6194	±0.0273
MMopt_C ₁	14			3.7729	±0.0166
MMopt	4.25			16.7924	±0.1145
	<i>2-Sample Design</i>				
MMopt_C ₁	1	13		2.1102	±0.0125
MMopt	1	9.5		2.2575	±0.0232
EDopt	1	24		2.6159	±0.0174
	<i>3-Sample Design</i>				
MMopt_C ₁	1	10.25	10.25	1.6967	±0.0078
MMopt	1	1	10.5	1.9991	±0.0192
EDopt	1	1	24	2.4194	±0.0174

- 1 Sample Design**: weighted MMopt performance approximates that of the weighted Bayesian optimal design (RMS error of 3.62 versus 3.77 AUC units)
- MMopt performance improves on EDopt design for 2 and 3 sample designs**
 - 2 Sample Design**: RMS error of 2.11 versus 2.62 (units of AUC)
 - 3 Sample Design**: RMS error of 1.70 versus 2.42 (units of AUC)
- All results are statistically significant to p<0.0001**

Summary

- **Multiple Model Optimal Design (MMOpt) provides an alternative approach to designing experiments**
 - Particularly attractive for Nonparametric Models (MM discrete prior)
 - Based on true MM formulation of the problem (i.e., classification theory)
 - Has several advantages relative to ED, EID and API (last year's PODE [23])
 - Based on recent theoretical overbound on Bayes Risk (Blackmore et. al. 2008 [4])
- **Introduced Weighted version of MMOpt which minimizes upper bound on the Weighted Bayes Risk**
 - Allows specification of costs for each type of classification error
 - Preserves overbound property so that weighted MMOpt designs are as straightforward to compute as unweighted MMOpt designs
 - Examples show that weighted MMOpt performance improves on EDopt, and compares favorably to the theoretically best performance of the weighted Bayes optimal classifier
- **MMOpt captures essential elements of Bayesian Experiment Design without the excessive computation**
 - Bayesian formulation of design problem for multiple model problems
 - Allows approximate pre-posterior analysis “without tears”
 - To be included in a future release of the USC *BestDose* software [3]

References (1/3)

- [1] Bayard D, Jelliffe R, Schumitzky A, Milman M, Van Guilder M. Precision drug dosage regimens using multiple model adaptive control: Theory and application to simulated Vancomycin therapy. In: *Selected Topics in Mathematical Physics, Prof. R. Vasudevan Memorial Volume*. Madras: World Scientific Publishing Co., 1995.
- [2] Schumitzky A. "Application of stochastic control theory to optimal design of dosage regimens," In: *Advanced methods of pharmacokinetic and pharmacodynamic systems analysis*. New York: Plenum Press; 1991:137-152
- [3] USC BestDose, <http://www.lapk.org>
- [4] Blackmore L, Rajamanoharan S, and Williams BC, "Active Estimation for Jump Markov Linear Systems," *IEEE Trans. Automatic Control.*, Vol. 53, No. 10., pp. 2223-2236, Nov. 2008.
- [5] Mallet A. "A maximum likelihood estimation method for random coefficient regression models," *Biometrika*. 1986;73:645-656.
- [6] B. Lindsay, "The Geometry of Mixture Likelihoods: a General Theory," *Ann. Statist.* 11: 86-94, 1983.
- [7] Schumitzky, "Nonparametric EM Algorithms for estimating prior distributions," *Applied Mathematics and Computation*, Vol. 45, Nol. 2, September 1991, Pages 143–157.

References (2/3)

- [8] R. Leary, R. Jelliffe, A. Schumitzky, and M. Van Guilder. "An Adaptive Grid Non- Parametric Approach to Pharmacokinetic and Dynamic (PK/PD) Population Models." In Computer-Based Medical Systems, 2001. CBMS 2001. Proceedings. 14th IEEE Symposium on, pp. 389-394. IEEE, 2001.
- [9] Y. Baek, "An Interior Point Approach to Constrained Nonparametric Mixture Models," Ph.D. Thesis, University of Washington, 2006.
- [10] Jelliffe R, Schumitzky A, Bayard D, Van Guilder M, Leary RH. "The USC*PACK Programs for Parametric and Nonparametric Population PK/PD Modeling," Population Analysis Group in Europe, Paris, France, June 2002.
- [11] D.Z. D'Argenio, A. Schumitzky, and X. Wang, ADAPT 5 User's Guide. Biomedical Simulation Resource, University of Southern California, 2009.
- [12] D'Argenio DZ, "Optimal Sampling Times for Pharmacokinetic Experiments," J. Pharmacokinetics and Biopharmaceutics, vol. 9, no. 6, 1981: 739-756
- [13] K. Chaloner and I. Verdinelli, "Bayesian experimental design: A review," Statistical Science, Vol. 10, No. 3, pp. 273-304, 1995.
- [14] L. Pronzato and E. Walter, "Robust experiment design via stochastic approximation," Mathematical Biosciences, Vol. 75, pp. 103-120, 1985

References (3/3)

- [15] E. Walter and L. Pronzato, "Optimal experiment design for nonlinear models subject to large prior uncertainties," *American Journal of Physiology – Regulatory, Integrative and Comparative Physiology*, 253:R530-R534, 1987.
- [16] M. Tod and J-M Rocchisani, "Comparison of ED, EID, and API criteria for the robust optimization of sampling times in pharmacokinetics," *J. Pharmacokinetics and Biopharmaceutics*, Vol. 25, No. 4, 1997.
- [17] D.Z. D'Argenio, "Incorporating prior parameter uncertainty in the design of sampling schedules for pharmacokinetic parameter estimation experiments," *Mathematical Biosciences*, Vol. 99, pp. 105-118, 1990.
- [18] Y. Merle and F. Mentre, "Bayesian design criteria: Computation, comparison, and application to a pharmacokinetic and a pharmacodynamic model," *J. Pharmacokinetics and Biopharmaceutics*, vol. 23, No. 1, 1995.
- [19] S.D. Silvey, *Optimal Design: An Introduction to the Theory for Parameter Estimation*. Chapman and Hall, London, 1980.
- [20] V.V. Fedorov, *Theory of Optimal Experiments*. Academic Press, New York, 1972.
- [21] R.O. Duda, P.E. Hart, D.G. Stork, *Pattern Classification*. John Wiley & Sons, New York, 2001.
- [22] L. Pronzato and A. Pazman, *Design of Experiments in Nonlinear Models*. Lecture Notes in Statistics, Springer, New York, 2013.
- [23] D.S. Bayard, R. Jelliffe and M. Neely, "Bayes Risk as an Alternative to Fisher Information in Determining Experiment Designs for Nonparametric Models," *Population Optimum Design of Experiments (PODE): Workshop*, Lilly UK, 15 June 2013.