

Optimal Designs for Individual Prediction in Population Based Models

Maryna Prus, Rainer Schwabe

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RCR model

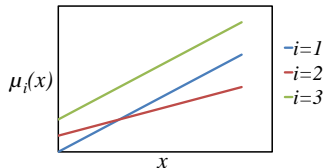
$$Y_{ij} = \mathbf{f}(x_{ij})^\top \boldsymbol{\beta}_i + \varepsilon_{ij}, \quad i = 1, \dots, n, \quad j = 1, \dots, m_i, \quad x_{ij} \in \mathcal{X}$$

- ▶ $\mathbf{f} = (f_1, \dots, f_p)^\top$
- ▶ $\boldsymbol{\beta}_i = (\beta_{i1}, \dots, \beta_{ip})^\top$, $E(\boldsymbol{\beta}_i) = \boldsymbol{\beta}$ unknown, $\text{Cov}(\boldsymbol{\beta}_i) = \sigma^2 \mathbf{D}$
- ▶ $E(\varepsilon_{ij}) = 0$, $\text{Var}(\varepsilon_{ij}) = \sigma^2$
- ▶ All $\boldsymbol{\beta}_i$ and all ε_{ij} uncorrelated

Gladitz & Pilz (1982): $\boldsymbol{\beta}$ known

Individual response:

$$\mu_i(x) = \mathbf{f}(x)^\top \boldsymbol{\beta}_i$$



Special model:

$$m_i = m \quad \& \quad x_{ij} = x_j$$

BLUE & BLUP in LMM

$$\mathbf{Y} = \mathbf{X}\boldsymbol{\beta} + \mathbf{Z}\boldsymbol{\gamma} + \boldsymbol{\varepsilon}$$

- ▶ $E(\boldsymbol{\gamma}) = 0$, $\text{Cov}(\boldsymbol{\gamma}) = \mathbf{G}$, \mathbf{G} regular
- ▶ $E(\boldsymbol{\varepsilon}) = 0$, $\text{Cov}(\boldsymbol{\varepsilon}) = \mathbf{R}$, \mathbf{R} regular
- ▶ $\boldsymbol{\gamma}$ and $\boldsymbol{\varepsilon}$ uncorrelated
- ▶ $\boldsymbol{\beta}$ unknown

$$\begin{pmatrix} \hat{\boldsymbol{\beta}} \\ \hat{\boldsymbol{\gamma}} \end{pmatrix} = \underbrace{\begin{pmatrix} \mathbf{X}^T \mathbf{R}^{-1} \mathbf{X} & \mathbf{X}^T \mathbf{R}^{-1} \mathbf{Z} \\ \mathbf{Z}^T \mathbf{R}^{-1} \mathbf{X} & \mathbf{Z}^T \mathbf{R}^{-1} \mathbf{Z} + \mathbf{G}^{-1} \end{pmatrix}^{-1}}_{=: \mathbf{C}} \begin{pmatrix} \mathbf{X}^T \mathbf{R}^{-1} \mathbf{Y} \\ \mathbf{Z}^T \mathbf{R}^{-1} \mathbf{Y} \end{pmatrix}$$

$$\mathbf{C} = \text{Cov} \begin{pmatrix} \hat{\boldsymbol{\beta}} \\ \hat{\boldsymbol{\gamma}} - \boldsymbol{\gamma} \end{pmatrix}$$

Henderson (1975)

Model adaptation I

▶ $Y_{ij} = \mathbf{f}(x_j)^\top \boldsymbol{\beta} + \mathbf{f}(x_j)^\top \mathbf{b}_i + \varepsilon_{ij}$

$$\mathbf{b}_i := \boldsymbol{\beta}_i - \boldsymbol{\beta}$$

▶ $\mathbf{Y}_i = \mathbf{F}\boldsymbol{\beta} + \mathbf{F}\mathbf{b}_i + \boldsymbol{\varepsilon}_i$

$$\mathbf{Y}_i = \begin{pmatrix} Y_{i1} \\ \vdots \\ Y_{im} \end{pmatrix} \quad \boldsymbol{\varepsilon}_i = \begin{pmatrix} \varepsilon_{i1} \\ \vdots \\ \varepsilon_{im} \end{pmatrix} \quad \mathbf{F} = \begin{pmatrix} \mathbf{f}(x_1)^\top \\ \vdots \\ \mathbf{f}(x_m)^\top \end{pmatrix}$$

▶ $\mathbf{Y} = (\mathbf{1}_n \otimes \mathbf{F})\boldsymbol{\beta} + (\mathbf{I}_n \otimes \mathbf{F})\mathbf{b} + \boldsymbol{\varepsilon}$

$$\mathbf{Y} = \begin{pmatrix} \mathbf{Y}_1 \\ \vdots \\ \mathbf{Y}_n \end{pmatrix} \quad \boldsymbol{\varepsilon} = \begin{pmatrix} \boldsymbol{\varepsilon}_1 \\ \vdots \\ \boldsymbol{\varepsilon}_n \end{pmatrix} \quad \mathbf{b} = \begin{pmatrix} \mathbf{b}_1 \\ \vdots \\ \mathbf{b}_n \end{pmatrix}$$

Problem: $\text{Cov}(\mathbf{b}) = \sigma^2 \mathbf{I}_n \otimes \mathbf{D}$ may be singular

Model adaptation II

$$\text{rank}(\mathbf{D}) = q, \quad q \leq p$$

$$\Rightarrow \exists \mathbf{H} \quad \text{with} \quad \mathbf{D} = \mathbf{H}\mathbf{H}^\top \quad \& \quad \text{rank}(\mathbf{H}) = q$$

$$\mathbf{c}_i := (\mathbf{H}^\top \mathbf{H})^{-1} \mathbf{H}^\top \mathbf{b}_i$$

$$\Rightarrow \mathbf{b}_i = \mathbf{H}\mathbf{c}_i \quad \& \quad \mathbb{E}(\mathbf{c}_i) = 0, \quad \text{Cov}(\mathbf{c}_i) = \sigma^2 \mathbf{I}_q$$

▶ $\mathbf{Y}_i = \mathbf{F}\boldsymbol{\beta} + \mathbf{F}\mathbf{H}\mathbf{c}_i + \varepsilon_i$

▶ $\mathbf{Y} = \underbrace{(\mathbf{1}_n \otimes \mathbf{F})}_{\mathbf{X}} \boldsymbol{\beta} + \underbrace{(\mathbf{I}_n \otimes (\mathbf{F}\mathbf{H}))}_{\mathbf{Z}} \mathbf{c} + \boldsymbol{\varepsilon} \quad \mathbf{c} = (\mathbf{c}_1^\top, \dots, \mathbf{c}_n^\top)^\top$

$$\Rightarrow \mathbf{R} = \sigma^2 \mathbf{I}_{nm} \quad \mathbf{G} = \sigma^2 \mathbf{I}_{nq}$$

Individual prediction

$$\begin{aligned}\hat{\beta}_i &= \hat{\beta} + \mathbf{H}\hat{c}_i \\ &= \mathbf{D}((\mathbf{F}^\top \mathbf{F})^{-1} + \mathbf{D})^{-1} \hat{\beta}_{i;\text{ind}} + (\mathbf{F}^\top \mathbf{F})^{-1}((\mathbf{F}^\top \mathbf{F})^{-1} + \mathbf{D})^{-1} \hat{\beta}\end{aligned}$$

$$\hat{\beta} = (\mathbf{F}^\top \mathbf{F})^{-1} \mathbf{F}^\top \bar{\mathbf{Y}} \quad \& \quad \hat{\beta}_{i;\text{ind}} := (\mathbf{F}^\top \mathbf{F})^{-1} \mathbf{F}^\top \mathbf{Y}_i$$

$$\bar{\mathbf{Y}} = \frac{1}{n} \sum_{i=1}^n \mathbf{Y}_i$$

MSE Matrix for predictors $(\hat{\beta}_1^\top, \dots, \hat{\beta}_n^\top)^\top$:

$$\sigma^2 \left(\left(\mathbf{I}_n - \frac{1}{n} \mathbf{1}_n \mathbf{1}_n^\top \right) \otimes (\mathbf{D} - \mathbf{D}((\mathbf{F}^\top \mathbf{F})^{-1} + \mathbf{D})^{-1} \mathbf{D}) + \frac{1}{n} \mathbf{1}_n \mathbf{1}_n^\top \otimes (\mathbf{F}^\top \mathbf{F})^{-1} \right)$$

for regular \mathbf{D} simplifies to

$$\sigma^2 \left(\left(\mathbf{I}_n - \frac{1}{n} \mathbf{1}_n \mathbf{1}_n^\top \right) \otimes (\mathbf{F}^\top \mathbf{F} + \mathbf{D}^{-1})^{-1} + \frac{1}{n} (\mathbf{1}_n \mathbf{1}_n^\top) \otimes (\mathbf{F}^\top \mathbf{F})^{-1} \right)$$

Individual design

Experimental settings x_1, \dots, x_m not necessarily all distinct

- ▶ Distinct settings: x_1, \dots, x_k
- ▶ Numbers of replications: m_1, \dots, m_k ; $\sum_{j=1}^k m_j = m$

Individual design:

$$\xi = \begin{pmatrix} x_1, \dots, x_k \\ w_1, \dots, w_k \end{pmatrix}$$

$$w_j = \frac{m_j}{m} \quad \& \quad \sum_{j=1}^k w_j = 1$$

Notation:

$$\mathbf{M}(\xi) := \sum_{j=1}^k w_j \mathbf{f}(x_j) \mathbf{f}(x_j)^\top = \frac{1}{m} \mathbf{F}^\top \mathbf{F} \quad \& \quad \mathbf{\Delta} := m\mathbf{D}$$

IMSE-criterion

$$\text{IMSE}_{\text{pred}}(\xi) := \int \mathbb{E}_{\xi} \left(\sum_{i=1}^n (\hat{\mu}_i(x) - \mu_i(x))^2 \right) \nu(dx)$$

$$\hat{\mu}_i(x) = \mathbf{f}(x)^\top \hat{\boldsymbol{\beta}}_i$$

$$\text{IMSE}_{\text{pred}}(\xi) = \frac{\sigma^2}{m} \text{tr} \left((\mathbf{M}(\xi)^{-1} + (n-1)(\boldsymbol{\Delta} - \boldsymbol{\Delta}(\mathbf{M}(\xi)^{-1} + \boldsymbol{\Delta})^{-1}\boldsymbol{\Delta})) \mathbf{V} \right)$$

$$\mathbf{V} := \int \mathbf{f}(x)\mathbf{f}(x)^\top \nu(dx)$$

For regular \mathbf{D} this simplifies to

$$\text{IMSE}_{\text{pred}}(\xi) = \frac{\sigma^2}{m} \text{tr} \left((\mathbf{M}(\xi)^{-1} + (n-1)(\mathbf{M}(\xi) + \boldsymbol{\Delta}^{-1})^{-1}) \mathbf{V} \right)$$

weighted sum of IMSE-criterion in fixed effects model and Bayesian IMSE

Optimality condition for approximate design

ξ^* *IMSE-optimal if and only if*

$$\begin{aligned} & (n-1) \mathbf{f}(x)^\top \mathbf{M}(\xi^*)^{-1} \left(\mathbf{M}(\xi^*)^{-1} + \mathbf{\Delta} \right)^{-1} \mathbf{\Delta} \mathbf{V} \mathbf{\Delta} \left(\mathbf{M}(\xi^*)^{-1} + \mathbf{\Delta} \right)^{-1} \mathbf{M}(\xi^*)^{-1} \mathbf{f}(x) \\ & \quad + \mathbf{f}(x)^\top \mathbf{M}(\xi^*)^{-1} \mathbf{V} \mathbf{M}(\xi^*)^{-1} \mathbf{f}(x) \\ & \leq \text{tr} \left(\left((n-1) \mathbf{\Delta} \left(\mathbf{M}(\xi^*)^{-1} + \mathbf{\Delta} \right)^{-1} \mathbf{M}(\xi^*)^{-1} \left(\mathbf{M}(\xi^*)^{-1} + \mathbf{\Delta} \right)^{-1} \mathbf{\Delta} + \mathbf{M}(\xi^*)^{-1} \right) \mathbf{V} \right) \end{aligned}$$

for all $x \in \mathcal{X}$

For regular \mathbf{D} this simplifies to

$$\begin{aligned} & \mathbf{f}(x)^\top \left((n-1) (\mathbf{M}(\xi^*) + \mathbf{\Delta}^{-1})^{-1} \mathbf{V} (\mathbf{M}(\xi^*) + \mathbf{\Delta}^{-1})^{-1} + \mathbf{M}(\xi^*)^{-1} \mathbf{V} \mathbf{M}(\xi^*)^{-1} \right) \mathbf{f}(x) \\ & \leq \text{tr} \left(\left((n-1) (\mathbf{M}(\xi^*) + \mathbf{\Delta}^{-1})^{-1} \mathbf{M}(\xi^*) (\mathbf{M}(\xi^*) + \mathbf{\Delta}^{-1})^{-1} + \mathbf{M}(\xi^*)^{-1} \right) \mathbf{V} \right) \end{aligned}$$

For any support point x_j of ξ^ equality holds*

Special case: random intercepts

- ▶ $f_1(x) \equiv 1$
- ▶ $\mathbf{D} = d_1 \mathbf{e}_1 \mathbf{e}_1^\top$, $\mathbf{e}_1 = (1, 0, \dots, 0)^\top$

Objective function:

$$\text{IMSE}_{\text{pred}}(\xi) = \sigma^2 \left(\frac{1}{m} \text{tr}(\mathbf{M}(\xi)^{-1} \mathbf{V}) + \frac{(n-1)d_1}{1 + m d_1} \nu(\mathcal{X}) \right)$$

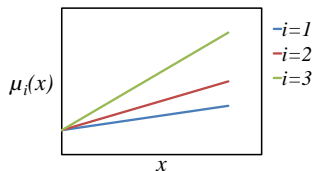
*IMSE-optimal design in fixed effects model
 remains IMSE-optimal for prediction*

Example: straight line regression

$$Y_{ij} = \beta_{i1} + \beta_{i2}x_j + \varepsilon_{ij}, \quad x_{ij} \in [0, 1]$$

Only slope is random

- ▶ $\beta_{i1} \equiv \beta_1$
- ▶ $\mathbf{D} = d_2 \mathbf{e}_2 \mathbf{e}_2^\top$

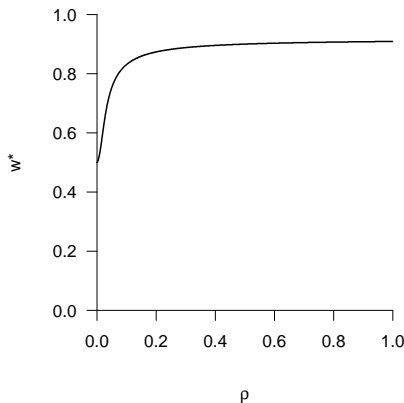


Individual design:

$$\xi^* = \begin{pmatrix} 0 & 1 \\ 1 - w^* & w^* \end{pmatrix}$$

Optimal weight

- ▶ $n = 100$
- ▶ $m = 10$
- ▶ $\rho = d_2/(1 + d_2)$

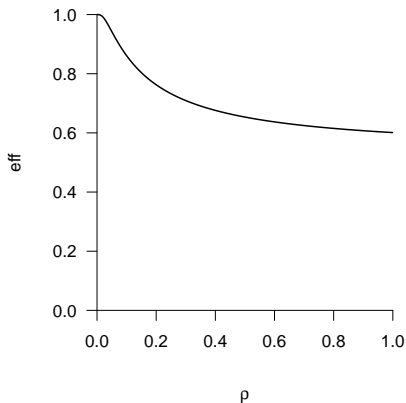


For small values of d_1

$$w^*(d_1) \approx w^* \quad \& \quad \text{Bayesian optimal design: } w_B^* = 1$$

Efficiency of equi-replicated design

- ▶ $n = 100$
- ▶ $m = 10$
- ▶ Equi-replicated design:
 $w = 0.5$



*Efficiency of Bayesian optimal design
for small values of d_1 : $eff(\xi_B^*) = 0$*

Outlook

- ▶ Exact design
- ▶ Different designs for different individuals
- ▶ Other design criteria
- ▶ ...

Thank you for your attention!

